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Abstract

 The drop size distribution (DSD) of precipitation particles is a fundamental property for characterizing rainfall. This study statistically clarified the characteristics of DSDs using approximately 10 years of DSD data obtained from a ground-based optical disdrometer in Kumagaya, the eastern part of Japan. The results showed that DSDs tended to maintain their shape even as rainfall intensity (*R*) increased, and they tended to be distributed in a 38 certain region, i.e., the mass-weighted mean diameter $(D_m) \sim 2.0-3.0$ mm and the 39 generalized intercept parameter (*N*_w) ~2×10³–3×10⁴ mm⁻¹ m⁻³, of the DSD parameter space defined by *D*^m and *N*w. The quasi-equilibrium shape of the DSDs, which is rarely observed only 16 cases in this study, was likely to be different characteristics between maritime and continental convection. Among them, the contribution to *R* was large when *D*^m or *N*^w was effectively increased with temporal change based on an error analysis. DSD characteristics were also identified by statistically evaluating the relationship between the specific 45 differential phase (K_{DP}) and *R* in DSDs for C-band polarimetric weather radar. The results 46 showed that the coefficient of the $K_{DP}-R$ relation tended to be larger (>24.0) during the warm season (from May to October) and smaller (<21.0) during the cold season (from January to April and from November to December) when assuming a temperature of 10 degrees Celsius, whereas the exponent of the relation had no apparent trend. Furthermore, it is likely that the slope parameter, one of the DSD parameters, can be optimized for stronger rainfall events with a nearly same size distribution.

52 **Keywords** raindrop size distribution; heavy rainfall; *K*_{DP}–*R* relation

1. Introduction

 The drop size distribution (DSD) of precipitation particles is a key property characterizing rainfall that has been studied worldwide in various ways, including by laboratory experiments, numerical modeling, and observational analyses. In cloud microphysical processes, the larger drops, once they grow to about 10 micrometers in radius by the condensation of water vapor, are produced by coalescence process before they reach the ground. The effects on raindrops of collisional coalescence and collisional breakup, which envisions to be obtained under warm rain processes, in particular have long been studied. The earlier studies showed that DSDs characteristically have shown to be trimodal (Valdez and Young 1985, Brown 1986, Feingold et al. 1988, Chen and Lamb 1994) based on Low and List (1982)'s parameterization and ground-based observations using impact type Joss-Waldvogel disdrometers, while laboratory experiments have shown them to be bimodal (Steiner and Waldvogel 1987; List et al. 1987; Asselin de Beauville et al. 1988). Subsequent studies showed that there were problems with the Joss-Waldvogel disdrometer showing a false peak mainly associated with uneven operation (Sheppard 1990) and this instrument-related peak was evaluated by McFarquhar and List (1993). McFarquhar (2004a) developed a more physical based parameterization of the fragment size distributions by colliding raindrops and showed that the stationary distribution obtained numerically was to be bimodal. More recently, Straub et al. (2010) also showed the bimodal stationary distribution obtained numerically using computational fluid dynamics program (Beheng et al. 2006; Schlottke et

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 al. 2010). These studies suggest that laboratory-based and theoretical parameterizations for numerically representing an equilibrium distribution are approaching closure (McFarquhar 2010). Nevertheless, observational studies (McFarquhar et al. 1996; D'Adderio et al. 2018) showed that DSDs rarely reached an equilibrium shape except in heavy rainfall events (Garsia-Garcia and Gonzalez 2000). Such observational results were not always consistent with the numerically obtained stationary distributions, probably due to evaporation (Brown 1993; Hu and Srivastava 1995), size sorting by updrafts (Kollias et al. 2001; McFarquhar and List 1991), number of observational samples (McFarquhar 2004b), and wind shear (Dawson II et al. 2015).

83 Although there were the inconsistencies between numerically obtained stationary 84 distributions and observed distributions as described above, a few studies tried to extract 85 stationary distributions from observational DSDs and to examine their characteristics. Low and List (1982) identified a relationship dependent on raindrop size to parameterize the 87 breakup efficiency in terms of collisional kinetic energy (CKE), and they showed that as a 88 result of raindrop coalescence, inflection points occur in DSDs at raindrop diameters of 89 around 1.0–2.6 mm. Thus, when collisional breakups occur, decreases in the frequency of raindrops with diameters around 1.5 mm and increases in the frequency of raindrops with diameters of less than 1 mm and with diameters of 2–3 mm often produce bimodal DSDs (McFarquhar 2004a; Prat and Barros 2007; McFarquhar 2010; Straub et al. 2010; D'Adderio et al. 2015) as stationary distributions, as shown in McFarquhar (2004a)'s Fig. 15 or Straub

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 observed during several field campaigns can be broadly classified into maritime and continental climatic regimes according to the DSD parameters (e.g., the mass-weighted mean diameter *D*^m and the generalized intercept parameter *N*w). Dolan et al. (2018) have conducted an empirical orthogonal function (EOF) analysis of global observations, including observations in different climate regions, and have proposed a region-independent DSD 119 regime that takes account of the relationship between the characteristics of DSDs in different 120 regions and their dominant cloud microphysical processes. It has been pointed out that in the tropics, in particular, it may not be possible to simply separate precipitation processes 122 below the melting layer into stratiform and convective rainfall regimes, which is considered 123 to be an important characteristic of precipitation processes that produce heavy rainfall in the tropics. For example, Thompson et al. (2018) showed that the convective-stratiform separation is not effective for rainfall estimation when *R* is stronger. In Japan, however, no study has demonstrated the characteristic features of DSDs based on long-term ground-based observations.

128 Quantitative precipitation estimation (QPE) of warm-season heavy rainfall remains a challenging task (e.g., Fritsch and Carbone 2004). *R* estimation techniques are the cornerstone of short-term (i.e., typically less than one hour) forecasting, and various relational formulas for *R* estimation have been proposed that use ground-based measurements and/or polarimetric weather radar observations (e.g., Ryzhkov and Zrnić 133 2019; WMO 2024). Among them, a method using the specific differential phase (K_{DP}) for

 is less affected by rainfall attenuation, is reasonable to use for C-band radar. Furthermore, to use operational polarimetric weather radar to accurately detect high-impact weather events that cause heavy rainfall, it is necessary to clarify the appropriate relation between 157 R and K_{DP}, which is expected to contain a signal that can be used for the detection of precursors of precipitating systems that may cause heavy rainfall-related disasters.

 Therefore, the first purpose of this study is to statistically characterize DSDs of strong rainfall events observed in Kumagaya, eastern Japan, and the second purpose is to obtain 161 an optimized K_{DP}-R relation for strong rainfall to improve QPE. In this paper, Section 2 presents details of the data and methodology, Section 3 presents the results, which are discussed in Section 4. In the final, conclusion section, some implications of this study are discussed.

2. Data and Methodology

 The data used were obtained by a first-generation OTT-Parsivel disdrometer (Löffler- Mang and Joss 2000) installed at Kumagaya (Fig. 1). A Parsivel disdrometer measures the 169 drop size and fall velocity of precipitation particles. The original data are observed within a laser beam sheet and output in 32 drop-size categories (*i* direction) and 32 fall-velocity categories (*j* direction). The laser beam sheet of the Parsivel disdrometer is 180 mm long, 172 30 mm wide, and 1 mm thick, and the effective sampling area *S* (m^{−2}) is expressed as 0.180 $173 \times (0.030 - L/2)$, where *L* is a size parameter that depends on the drop size. The drop

Fig. 1

174 diameters range from 0.2 to 25 mm and velocities range from 0.2 to 20 m s⁻¹, respectively. 175 The two smallest drop size categories were not used in this study because of their low signal-176 to-noise ratios (Tokay et al. 2013). Observed drop sizes were converted into the equivalent 177 volume diameter (*D*e) in accordance with Adachi et al. (2013) (hereafter, *D*^e is represented simply as *D*). The drop-size distribution *N*(*D*) (mm−1 m−3 178) was then obtained as follows:

179
$$
N(D) = \sum_{j=1}^{32} \frac{C_{i,j}}{S V_j(D_i) \, \Delta t \, \Delta D_i}, \quad (1)
$$

180 where *Ci,j* represents the raw output having a matrix structure with 32 drop-size categories 181 and 32 fall-velocity categories, $V_j(D_i)$ (m s⁻¹) is the fall velocity for each drop size *D_i*, ∆*t* is the 182 observation time interval (60 s in this study), ΔD_i is the drop-size spread (mm) in equivalent 183 volume diameters, and the subscripts *i* and *j* indicate the drop size and fall velocity category, 184 respectively. To eliminate unreliable data, disdrometer observations with *R* less than 0.1 mm 185 hr^{−1}, or a number of observed particles less than 10 drops per 1 min in *C_{i,j}*, was excluded 186 following Tokay et al. (2014). We verified *R* data against a tipping-bucket-type rain gauge 187 data. First, we assumed that the formula of Atlas et al. (1973) as described the fall velocity 188 of raindrops and used data with fall velocities within ±50% of the formula values. After this 189 quality control step, *R* (mm hr^{−1}) was calculated from the observed DSD data with the 190 following equation:

191
$$
R = 6 \times 10^{-4} \pi \sum_{i=1}^{32} \sum_{j=1}^{32} \frac{C_{i,j}}{S V_j(D_i) \Delta t \Delta D_i} V_j(D_i) D_i^3, (2)
$$

192 which is the mass flux, i.e., the mass of raindrops falling on a unit area per unit time.

193 Figure 2 shows a scatter plot of *R* calculated with Eq. (2) using DSDs observed by the Fig. 2

 Parsivel disdrometer versus those obtained by the rain gauge. The values generally fell along a straight one-to-one line (correlation coefficient 0.92; slope of the fitted line 0.95); this result indicated that the Parsivel disdrometer-related *R* were good agreement with rain gauge data.

198 The DSD parameters were calculated by assuming a normalized gamma distribution and 199 calculated the parameter using the moment technique (Hardin and Guy 2017). The 200 normalized gamma distribution of drop sizes (Testud et al. 2001) was expressed as

$$
N(D) = N_w f(\mu) \left(\frac{D}{D_m}\right)^{\mu} \exp\left[-(4+\mu)\frac{D}{D_m}\right], \quad (3)
$$

202 where the generalized intercept parameter N_w (mm⁻¹ m⁻³) was calculated as

203
$$
N_{\rm w} = \frac{4^4 \, LWC}{\pi \rho_{\rm w} \, D_{\rm m}^4}, \quad (4)
$$

204 and $f(\mu)$ was defined as

205
$$
f(\mu) = \frac{\Gamma(4)}{4^4} \frac{(4+\mu)^{(4+\mu)}}{\Gamma(4+\mu)}.
$$
 (5)

206 Here, μ is a shape parameter, ρ_w is the density of water (g m⁻³), and Γ is the Gamma 207 function. The liquid water content *LWC* (g m⁻³) and the mass-weighted mean diameter *D*_m 208 (mm) were obtained with following equations:

209
$$
LWC = \frac{\pi \rho_w}{6} \sum_{i=1}^{32} \sum_{j=1}^{32} \frac{C_{i,j}}{S V_j(D_i) \Delta t \Delta D_i} D_i^3, \quad (6)
$$

210 and

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211
$$
D_{\rm m} = \frac{\sum_{i=1}^{32} \sum_{j=1}^{32} \frac{C_{i,j}}{S V_j(D_i) \Delta t \Delta D_i} D_i^4}{\sum_{i=1}^{32} \sum_{j=1}^{32} \frac{C_{i,j}}{S V_j(D_i) \Delta t \Delta D_i} D_i^3}.
$$
 (7)

212 The exponential distribution of drop size was expressed as

$$
N(D) = N_0 \exp(-\Lambda D), \quad (8)
$$

214 where N_0 (mm⁻¹ m⁻³) is the intercept parameter and Λ (mm⁻¹) is the slope parameter, 215 which are calculated by the moment relation method of Zhang et al. (2008).

 The analysis period was from 00:00 Japan Standard Time (JST: 9 hours ahead of Coordinated Universal Time) on 1 January 2010 to 23:59 JST on 31 December 2022. The data were output every 1 min. Several outages for instrument maintenance occurred during the analysis period; the actual number of available samples was 294,756.

 To detect the equilibrium shape of DSDs, the algorithm of D'Adderio et al. (2015, 2018) 221 was applied to Parsivel disdrometer data. The algorithm is based on the concept that a DSD that has an inflection point caused by a decrease in the frequency of raindrops with a diameter of around 1.5 mm, and an increase in the frequency of larger and smaller drops as a result of collisional coalescence and breakup. Briefly, the calculation procedure of D'Adderio et al. (2015) is as follows: Over 5 bins from smaller (i.e., 1.0 mm in their study and we used the same value in this study) to larger diameters, the linear best fit of the 227 considered DSD obtained from Eq. (1) is calculated. Starting points are considered from 1.0 to 1.6 mm with 0.2-mm-diameter spreads, and then four linear relationships are calculated. 229 The maximum slope among the four relationships is defined as the highest slope (HS; units,

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3. Results

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 shows an inflection point in the distributions is examined (Fig. 5). The size ranges from 1.0 to 1.6 mm with the most frequent diameter of 1.59 mm, indicating that the shape of DSDs seems to be approaching an equilibrium shape although there are small variations in the inflection points of the distribution.

Fig. 5

 We next examined the contributions of temporal changes in *D*^m and *N*^w to changes in *R* in order to evaluate how *R* increases or decreases in terms of DSD parameters. Additionally, we applied error analysis (e.g., Bevington and Robinson 2003) to quantify the contribution of changes in *D*^m and *N*^w to changes in *LWC*. First, we transformed Eq. (4) to the following:

288
$$
LWC = \frac{\pi \, \rho_w}{4^4} \, N_w \, D_m^4, \quad (9)
$$

289 where $\frac{\pi \rho_w}{4^4}$ is a constant. Next, taking the logarithm of both sides of Eq. (9) and

Fig. 6

differentiating each side, we obtained the following equation:

$$
\frac{\Delta LWC}{LWC} = \frac{\Delta N_{\rm w}}{N_{\rm w}} + 4 \frac{\Delta D_{\rm m}}{D_{\rm m}}.\tag{10}
$$

 Here, ∆ indicates the difference in parameter values between the time when HS > 0 and one minute before. Note that Eqs. (2) and (6) can be regarded as approximately proportional to 294 the integral of the DSD multiplied by D to the $3rd$ moments; thus, the analysis in this study was performed on an assumption that the change in *R* was proportional to that in *LWC*. Here,

we focused on QEDSD cases and applied the assumption to the analysis.

297 A scatter plot between the two terms on the right-hand side of Eq. (10) in relation to the change of *R* for QEDSDs is shown in Fig. 6. The QEDSD cases of increasing *R* were distributed in the first (star), second (square), and fourth (triangle) quadrants, whereas those cases of decreasing *R* were less likely to be located in the first quadrant. Comparison of the frequencies of cases among the quadrants indicated that the highest number of cases (8 DSDs) were those in which *N*^w increased and *D*^m decreased, followed by cases in which both *N*^w and *D*^m increased (4 DSDs), and those in which *D*^m increased but *N*^w decreased (3 DSDs). Note that in 3 DSDs among the highest number of cases rainfall amounts rapidly increased. Only 1 DSDs showed decreases in both *N*^w and *D*m. Compared to Fig. 3, half of the second quadrant cases with increased *N*^w and decreased *D*^m were distributed to correspond to the cases with larger *D*^m (>3.0 mm) cases, which is mainly related to the wide range of frequency distributions in the shape of QEDSDs (Fig. 4b). The other half of second (*D*^m decreased and *N*^w increased) and most of the fourth (*D*^m increased and *N*^w decreased)

- 310 quadrant cases were mostly distributed between maritime and continental convective areas
- 311 except the first (both *D*^m and *N*^w increased).
- 312 When DSDs approach an equilibrium shape, Λ approaches a constant value (e.g., Willis 313 and Tattelman 1989) as described in the introduction. Therefore, we investigated the 314 relationship between *R* and Λ (Fig. 7). The DSDs displayed a typical shape such that for larger value of *R*, Λ values were concentrated ranging from 1.4 to 2.0 mm⁻¹; this trend was 316 especially noticeable for stronger *R* (e.g., >10 mm hr⁻¹). In addition, the QEDSD cases (red 317 crosses) were mainly distributed in the area with Λ < 2 mm⁻¹, which is consistent with the 318 previous studies (e.g., Fredrich et al. 2013; Unuma 2024). Fig. 7
- 319 The K_{DP}–*R* relation used to estimate *R* in DSDs was also examined in the long-term 320 disdrometer data because an optimal $K_{DP}-R$ relation, which is less dependent on DSD 321 (Sachidananda and Zrnić 1987) than the other relations, is expected to be obtained. The 322 *K*_{DP}–*R* relation calculated was as follows:
-

$$
R(K_{\rm DP}) = a K_{\rm DP}^b, \quad (11)
$$

 where coefficient *a* and exponent *b* are constant. The coefficients (Table 1) were estimated 325 from ground-based DSD data at Kumagaya during each month, where *K*_{DP} was obtained from the Parsivel disdrometer data by the *T*-matrix method (Mishchenko et al. 1996) using a drop-shape relationship proposed by Beard and Chuang (1987), assuming the temperature of 10 degrees C, and the transmitting frequency of 5.4 GHz in the temporal interval of 1 min, and then *R* was obtained by Eq. (2). In general, the values of *a* and *b* at Table 1

 Kumagaya are smaller than those shown previously (Yamauchi et al. 2012; Adachi et al. 2015). The value of *a* tends to be larger in the warm (e.g., from May to October) season and smaller in the cold (e.g., from January to April and from November to December) season (Table 1), whereas *b* shows a large monthly variation but no apparent trend.

4. Discussion

4.1 Quasi-Equilibrium DSD obtained in this study

 QEDSDs which resemble numerically obtained stationary distribution are likely to occur in Japan. Previous studies have mainly discussed the equilibrium shape of DSD based on observations in tropical and mid-latitude regions (e.g., Zawadzki and De Agostinho Antonio 1988; McFarquhar et al. 1996; D'Adderio et al. 2018). Recently, it has been pointed out that the equilibrium shape of DSD can also be observed in subtropical regions for the *R* >15 mm 342 hr⁻¹ condition (Murata et al. 2020). In addition, equilibrium shape of DSD was observed in a very humid environment when heavy rainfall was brought in Japan (Misumi et al. 2021; Unuma et al. 2023; Unuma 2024). These observations and the characteristics obtained in this study suggest that equilibrium shape of DSD can also be observed in a warm and humid climate region such as Japan. Although the environmental variability (i.e., thermodynamic properties and/or vertical wind shear) may affect the DSD shape (e.g., Hu and Srivastava 1995; Kollias et al. 2001; McFarquhar and List 1991; Dawson II et al. 2015), such effects need to be investigated and are left for a future work.

 As shown in Fig. 3, QEDSD cases were not associated with the specific parameter space, 368 i.e., stronger *R* with *D*_m ranging from 1.5 to 3.0 mm and *N*w >3×10³ mm⁻¹ m⁻³. Thus, the 169 characteristics with the conditions of *R* > = 40 mm hr⁻¹ and Λ < 2 mm⁻¹ based on both Figs.

 $370\quad \,$ 4 and 7 were examined. Figure 8 shows normalized DSDs as well as Fig. 4. A shape of the \mid Fig. 8 distributions is quite similar to the all data (Fig. 4a), but significantly different in the number concentrations between 1.5 mm and 3.0 mm in diameter, and the percentile ranges are narrower in Fig. 8 than in Fig. 4a. In addition, the distributions are like an exponential distribution except lower concentrations in smaller (<0.5 mm) diameter, which means that 375 nearly the same distribution could be assumed in terms of exponential distribution. The K_{DP}-*R* relation for stronger *R* will be discussed in the following subsection.

4.2 Toward improvement of QPE in Japan

378 Previous studies have shown that the coefficient *a* and exponent *b* of the *K*_{DP}–*R* relation vary regionally (Table 2). The coefficients obtained from C-band polarimetric radar and/or ground-based observations at different locations showed that the value of coefficient *a* (25.1) at Kumagaya was close to values in the UK (24.7) and Okinawa, Japan (28.8), whereas the value of exponent *b* at Kumagaya was close to values in Tropics (0.78) and Taiwan (0.80). In contrast, the value of the coefficient *a* at Tokyo (30.9) was larger than its value at Kumagaya (25.1), probably due to the vertical variabilities of radar reflectivity (Zawadzki 1984; Austin 1987; Kitchen and Jackson 1993; Anagnostou and Krajewski 1999; Villarini and Krajewski 2010). For example, the Japan Meteorological Agency uses ground-based observations (i.e., rain gauge data) to adjust *R* values obtained from weather radar observations at an altitude of roughly 2 km (Nagata 2011), even if *R* changes in the vertical direction. However, vertical differences in *R* need to be considered when weather radar is Table 2 being used to capture conditions before heavy rainfall is observed at ground level, and they are also important for capturing precursors to heavy rainfall. An investigation of this point was outside the scope of this study but remains for future work.

393 The results of this study suggested that the *K*_{DP}–*R* relation used in Yamauchi et al. (2012) and Adachi et al. (2015) for C-band polarimetric weather radar observations could be improved for strong rainfall events by considering the condition with $Λ < 2$ mm⁻¹ because 396 their studies were based on one event and did not examine the *K*_{DP}–*R* relation directory using in situ observational data. Yamauchi et al. (2012) showed that the estimated *R* using 398 the K_{DP}–*R* relation proposed by Bringi and Chandrasekar (2001) was in good agreement with the *R* obtained from an optical disdrometer for only one heavy rainfall event occurred 400 in August. The K_{DP}-R relation used in their study is based on weather radar related coefficients (i.e., not near the ground but above a few km), which may cause a little higher 402 value of coefficient *a* as shown in Table 2. Adachi et al. (2015) also used the same $K_{DP}-R$ relation and compared *R* for only one event that occurred in December. As described in the 404 previous subsection, the K_{DP}-R relation for stronger R was examined. The obtained coefficients were quite similar to the annual mean value, but the value of *a* is a little larger in the cases of stronger *R* cases than in the annual mean. These coefficients were probably optimized for heavy rainfall events because nearly the same size distribution in terms of the slope parameter was obtained (Fig. 8 and Table 1). Consequently, the slope parameters obtained in this study could be used to optimize the *KDP*–*R* relation, which has been

 specifically improved for Japan, to make a QPE method suitable for heavy rainfall in operational use.

5. Summary and Conclusions

 The characteristics of DSDs were statistically clarified by using about 10 years of DSD data obtained in Kumagaya, eastern Japan. The results showed that DSDs tended to be 416 distributed in a specific area, i.e., *D*_m ~2.0–3.0 mm and N_w ~2×10³–3×10⁴ mm⁻¹ m⁻³, in the 117 DSD parameter (D_m–N_w) space when *R* is stronger (>=40 mm hr⁻¹). For the quasi- equilibrium DSDs, the contribution of temporal changes in both *N*^w and *D*^m to increments of *R* was found to be large whereas the contribution to *R* was large when *D*^m or *N*^w was effectively increased with temporal change based on an error analysis.

421 In terms of QPE, the coefficient and exponent of the $K_{DP}-R$ relation, which had previously been calculated on a case-by-case basis, were statistically evaluated to identify characteristics that depended on seasonality and DSD variation. The results showed that 424 the coefficient in the *K*_{DP}–*R* relation tended to be larger in the warm season (from May to 425 October) and smaller in the cold season (from January to April and from November to December) when assuming the fixed temperature of 10 degrees Celsius, whereas the 427 exponent of the relation had no apparent trend. It is likely, however, that the *K*_{DP}–*R* relation 428 could be optimized for heavy rainfall events in Japan by using the slope parameter.

The present results should be statistically evaluated by using two-dimensional

 (geospatial) distributions or three-dimensional (geospatial and vertical) information of the three-dimensional structures of precipitating systems, obtained from polarimetric weather radar observations. The present results, which provided a basis for such a statistical evaluation for the first time, are expected to be very useful in this regard. An evaluation using polarimetric weather radar observations, i.e., vertical variations of DSDs in terms of process- oriented perspective, was outside the scope of this study and remains for future work. To improve understanding of the characteristics of DSDs and cloud microphysical processes based on observational facts, analyses such as the present study should be conducted in multiple regions in different climatic zones.

Data Availability Statement

 The research data and code used in this study are available from the corresponding author on request.

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Appendix

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- 740 stationary distributions, $N(D) = 4 \times 10^4$ exp(−ΛD) with Λ = 2.0 mm⁻¹ (black dash) and 1.0
- $741 \qquad \text{mm}^{-1}$ (black dot) are displayed as a reference.

Fig. 1. Location of the in situ precipitation measurement site in Kumagaya, Japan (cross), where both a first-generation OTT-Parsivel disdrometer (upper left) and a tipping-bucket-type rain gauge are installed at the Kumagaya Local Meteorological Office, Japan Meteorological Agency.

Fig. 2. Scatter plot of the first-generation OTT-Parsivel data versus tippingbucket-type rain gauge data in relation to the 1-minute rainfall intensity (mm hr⁻¹) (color scale). The equation of the line fitted to the data (blue line) and the correlation coefficient between the two datasets are displayed in the upper left.

Fig. 3. Scatter plot of the mass-weighted mean diameter (*D*m; mm) versus the generalized intercept parameter (*N*w; mm[−]¹ m[−]³) in 1-min drop size distribution observations in relation to rainfall intensity (*R*; mm hr[−]¹) (color scale). Regions in the two-dimensional space of continental convection (solid line rectangle), maritime convection (dashed line rectangle), and stratiform conditions (dotted line), as classified by Bringi et al. (2003), are also shown. The markers of star, square, cross, and triangle represents the data of the first, second, third, and fourth quadrants categorized in Fig. 6, respectively.

(a) Raindrop Size Distribution at Kumagaya # of DSDs = 294756 10^o 10^{-1} $N(D)/N_t$ (mm⁻¹) 10^{-2} 10^{-3} 10^{-4} $\frac{1}{2}$ ӟ $\dot{0}$ $\mathbf 1$ 4 5 D_e (mm) (b) Raindrop Size Distribution at Kumagaya # of $DSDs = 16$ 10^o McFarguhar (2004) Straub et al. (2010) 10^{-1} $N(D)/N_t$ (mm⁻¹) 10^{-2} 10^{-3} 10^{-4} $\dot{0}$ i $\frac{1}{2}$ $\frac{1}{3}$ $\dot{4}$ 5 $D_{\rm e}$ (mm)

Fig. 4. Drop size distributions (*N*(*D*); mm[−]¹ m[−]³) normalized by the total number concentrations during each 1-min period (*N*t; m[−]³): (a) all 1-min data, and (b) 1 min data with highest slope (HS; mm⁻² m⁻³) > 0 and coefficient of determination regression score of the two stationary distributions (McFarquhar 2004a or Straub et al. 2010) equal to or larger than 0.9. The reference distributions obtained by McFarquhar (2004a), and Straub et al. (2010) are also shown. The black solid lines show median values, the gray dashed lines show the 25th and $75th$ percentile values, and the gray dotted lines show the 10th and 90th percentile values.

Fig. 5. Frequency distribution of equivalent volume diameter (*D*e; mm) at inflection point occurred corresponding to the quasi-equilibrium drop size distributions. The *D*^e categories are 1.05, 1.17, 1.35, and 1.59.

Fig. 6. Scatter plot between the two terms on the right-hand side of Eq. (10) in relation to differences of rainfall intensity (*R*) between the time when a quasiequilibrium drop size distributions was detected, and 1 min earlier (color scale). The data of the first, second, third, and fourth quadrants were categorized with the markers of star, square, cross, and triangle, respectively, as well as Fig. 3.

Fig. 7. Scatter plot of slope parameter (Λ; mm[−]¹) versus rainfall intensity (*R*; mm hr^{−1}) in relation to the two-dimensional frequency of occurrence (color scale). The crosses indicate the quasi-equilibrium drop size distribution cases.

Fig. 8. The same as Fig. 4, but for rainfall intensity *R* >= 40 mm hr[−]¹ and the slope parameter Λ < 2 mm⁻¹.

Fig. A1. The stationary distributions for McFarquhar (2004a) (blue) and Straub et al. (2010) (orange). To compare the slope parameter of the exponential distribution with the stationary distributions, $N(D) = 4 \times 10^4$ exp($\neg D$) with Λ = 2.0 mm⁻¹ (black dash) and 1.0 mm⁻¹ (black dot) are displayed as a reference.

Table 1: Coefficient *a* and exponent *b* of the *K*_{DP}–*R* relation (Eq. (11)) obtained from this study.

Table 2: Values of coefficient *a* and exponent *b* obtained in various locations: Italy (Scarchilli et al. 1993; Vulpiani et al. 2012), Australia (May et al. 1999; Keenan et al. 2001), Japan (Yamauchi et al. 2012; This study), UK (Bringi et al. 2011), France (Figueras et al. 2012), Taiwan (Wang et al. 2013), and the tropics (Thompson et al. 2018). Also shown is the transmitting frequency of the radar *f* (GHz) used in each location. Note that, Yamauchi et al. (2012)'s coefficients are based on Bringi and Chandrasekar (2001).

